## 12 Bivariate Data Analysis: Regression and Correlation Methods

### 12.1 Introduction (P.187-191)

Many scientific investigations often involve two continuous variables and researchers are interested to know whether there is a (linear) relationship between the two variables. For example, a researcher wishes to investigate whether there is a relationship between the age and the blood pressure of people 50 years or older.

A Motivational Example: Suppose that a medical researcher wishes to investigate whether different dosages of a new drug affect the duration of relief from a particular allergic symptoms.

To study this, an experiment is conducted using a random sample of ten patients and the following observations are recorded:

| Dosage $(x)$ | 3 | 3 | 4 | 5 | 6 | 6 | 7 | 8 | 8 | 9 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration of relief $(y)$ | 9 | 5 | 12 | 9 | 14 | 16 | 22 | 18 | 24 | 22 |

Note: There are two variables in the problem and they are labelled by $x$ and $y$ for our convenience. The common sense suggests that the variable $y$ depends on $x$ and $x$ can be independently selected and controlled by the researcher. However, the variable $y$ cannot be controlled and is dependent on $x$.

## Further Examples

| Variable $1(x)$ | Variable 2 $(y)$ |
| :--- | :--- |
| Income | Expenditure |
| Temperature | Reaction time of a chemical |
| Alcohol consumption | Cholesterol level |
| Age of cancer patients | Length of survival |

In these cases the variable $y$ may depend on $x$. Therefore, $x$ can be considered as an independent variable and the variable $y$ as dependent (on $x$ ) variable. Statisticians are often interested to see whether there is a linear relationship (or linear association) between the two variables $x$ and $y$. Such observations as collected as pairs on $x$ and $y$ (or $(x, y)$ ) are called bivariate data.

In the previous bivariate example, ' $x_{1}=3$ ' corresponds to ' $y_{1}=$ 9 ' and there are $n=10$ pairs.

## A Notation

In general, there are $n$ pairs of such bivariate data given by

$$
\left(x_{1}, y_{1}\right),\left(x_{2}, y_{2}\right), \ldots,\left(x_{n}, y_{n}\right)
$$

### 12.2 Graphical Representation of Bivariate Data

A standard plot on a grid paper of $y$ (y-axis) against $x$ ( x -axis) gives a very good indication of the behaviour of data. This coordinate plot of the points $\left(x_{1}, y_{1}\right),\left(x_{2}, y_{2}\right), \ldots,\left(x_{n}, y_{n}\right)$ on a grid paper is called a scatter plot.

Note: The first step of the analysis of bivariate data is to plot the observed pairs, $(x, y)$ and obtain a scatter plot. This plot
gives a clear picture of a possible relationship between $x$ and $y$. Now we look at a number of other possible scatter plots we may observe in data analysis.

## Some typical scatter plots

(a) Positive slope (or upward trend)

(b) Negative slope (or downward trend)

(c) Random scatter (or no apparant pattern)


Example: A scatter plot of the above bivariate data is:


This scatter plot shows that the points seem to cluster around a straight line. It tells us that there is a possible (approximate) liner relationship between $x$ and $y$.

Then we want to further investigate:
Is the linear relationship between the variables $x$ and $y$ clear and significant

We need to find a measure to investigate the strength of a possible linear relationship between two variables $x$ and $y$. This measure is known as the correlation coefficient (or Pearson's correlation coefficient) between $x$ and $y$.

### 12.3 The correlation coefficient (P.215-218)

Recall the following:

$$
\begin{gathered}
L_{x y}=\sum_{i=1}^{n} x_{i} y_{i}-\frac{1}{n}\left(\sum_{i=1}^{n} x_{i}\right)\left(\sum_{i=1}^{n} y_{i}\right), \\
L_{x x}=\sum_{i=1}^{n} x_{i}^{2}-\frac{1}{n}\left(\sum_{i=1}^{n} x_{i}\right)^{2}
\end{gathered}
$$

and

$$
L_{y y}=\sum_{i=1}^{n} y_{i}^{2}-\frac{1}{n}\left(\sum_{i=1}^{n} y_{i}\right)^{2}
$$

( $L_{x x}$ is used in the calculation of $s^{2}$ ). Now the correlation coefficient between $x$ and $y$ is denoted by $r$ and is given by the ratio

$$
r=\frac{L_{x y}}{\sqrt{L_{x x} L_{y y}}} .
$$

Important Property: It can be shown that the values of $r$ lie between -1 and 1 , or

$$
-1 \leq r \leq 1
$$

for all such calculations. When $r$ is large and close to +1 or -1 (ideally $0.5<r<1$ or $-1<r<-0.5$ ), we conclude that there is a linear relationship between $x$ and $y$.
Note: All these formulae are available on the formulae sheet.
Read: P.222-223.

Example: Find the correlation coefficient, $r$, between $x$ and $y$ for the data in the dosage example.

| Dosage $(x)$ | 3 | 3 | 4 | 5 | 6 | 6 | 7 | 8 | 8 | 9 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Duration of relief $(y)$ | 9 | 5 | 12 | 9 | 14 | 16 | 22 | 18 | 24 | 22 |

## Solution:

It is easy to obtain that:

$$
\begin{gathered}
\sum_{i=1}^{10} x_{i}=59, \sum_{i=1}^{10} x_{i}^{2}=389 \\
\sum_{i=1}^{10} y_{i}=151, \sum_{i=1}^{10} y_{i}^{2}=2651, \sum_{i=1}^{10} x_{i} y_{i}=1003
\end{gathered}
$$

Now,

$$
\begin{aligned}
L_{x x} & =\sum_{i} x_{i}^{2}-\frac{\left(\sum_{i} x_{i}\right)^{2}}{n}= \\
L_{y y} & =\sum_{i} y_{i}^{2}-\frac{\left(\sum_{i} y_{i}\right)^{2}}{n}=
\end{aligned}
$$

and

$$
L_{x y}=\sum_{i} x_{i} y_{i}-\frac{\left(\sum_{i} x_{i}\right)\left(\sum_{i} y_{i}\right)}{n}=
$$

Hence

$$
r=\frac{L_{x y}}{\sqrt{L_{x x} L_{y y}}}=
$$

This correlation coefficient is close to one. From the scatter plot, we've noticed that the points are very close to a straight line with a positive slope.

## In general:

- It can be seen that when the points lie perfectly on a straight a line, then $r$ is either +1 (when the slope is positive) or -1 (when the slope is negative).
- If the points are close to a straight line (not perfectly), then $r$ is close to either 1 or -1 (but not exactly equal).


## Illustrative Examples:

Find the correlation coefficient between $x$ and $y$ for the following two data sets:

1. | $x$ | 0 | 1 | 2 | 3 |
| :---: | :---: | :---: | :---: | :---: |
| $y$ | 3 | 8 | 13 | 18 |

Note that here $y=5 x+3$.

2. | $x$ | 0 | 1 | 2 | 3 |
| :--- | :--- | :--- | :--- | :--- |
| $y$ | 8 | 6 | 4 | 2 |

Note that here $y=-2 x+8$.

Solution: We have

1. $\sum_{i} x_{i}=6, \sum_{i} x_{i}^{2}=14, \sum_{i} y_{i}=42, \sum_{i} y_{i}^{2}=566, \sum_{i} x_{i} y_{i}=88$

$$
\begin{aligned}
& L_{x x}=\sum_{i} x_{i}^{2}-\frac{\left(\sum_{i} x_{i}\right)^{2}}{n}= \\
& L_{y y}=\sum_{i} y_{i}^{2}-\frac{\left(\sum_{i} y_{i}\right)^{2}}{n}=
\end{aligned}
$$

$\qquad$

$$
L_{x y}=\sum_{i} x_{i} y_{i}-\frac{\left(\sum_{i} x_{i}\right)\left(\sum_{i} y_{i}\right)}{n}=
$$

and

$$
r=\frac{L_{x y}}{\sqrt{L_{x x} L_{y y}}}=
$$

$\qquad$
2. $\sum_{i} x_{i}=6, \sum_{i} x_{i}^{2}=14, \sum_{i} y_{i}=20, \sum_{i} y_{i}^{2}=120, \sum_{i} x_{i} y_{i}=20$

$$
\begin{gathered}
L_{x x}=\sum_{i} x_{i}^{2}-\frac{\left(\sum_{i} x_{i}\right)^{2}}{n}=\underline{14-6^{2} / 4=5} \\
L_{y y}=\sum_{i} y_{i}^{2}-\frac{\left(\sum_{i} y_{i}\right)^{2}}{n}=\underline{120-20^{2} / 4=20} \\
L_{x y}=\sum_{i} x_{i} y_{i}-\frac{\left(\sum_{i} x_{i}\right)\left(\sum_{i} y_{i}\right)}{n}=\underline{20-6(20) / 4=-10}
\end{gathered}
$$

and

$$
r=\frac{L_{x y}}{\sqrt{L_{x x} L_{y y}}}=
$$

$\qquad$

Exercise: Draw a scatter plot for each of the examples.

## Notes:

1. Using a similar argument, it can be seen that when $r$ is close to +1 or -1 , the points are very close to a straight line as given below:

Case I: $r \approx+1$ :


In this case $r$ is close to +1 and we say that there is a strong positive linear relationship between $x$ and $y$.

Case II: $r \approx-1$ :


In this case $r$ is close to -1 and we say that there is a strong negative linear relationship between $x$ and $y$.
2. When $r$ is close to zero (0), the points are far away from a straight line as shown below:


In this case we say that there is a very weak (or no) linear relationship between $x$ and $y$ since the points are randomly scattered.

## A Linear Relationship between $x$ and $y$

When $r$ is close to +1 or -1 , there is a possible linear relationship between $x$ and $y$ as described by $y=a+b x$. In such cases, it can be shown that $r^{2}$ gives the proportion of variability explained by a linear relationship between $x$ and $y$.


In other words, it is 'one minus the proportion of variation not explained by the model' as given below:

$$
r^{2}=1-\frac{\sum_{i} r_{i}^{2}}{\sum_{i} d_{i}^{2}}=1-\frac{\sum_{i}\left(y_{i}-\hat{y}_{i}\right)^{2}}{\sum_{i}\left(y_{i}-\bar{y}\right)^{2}}
$$

Hence it measures an overall fit of the regression model.
Note: Linear relationship between $x$ and $y$ should be investigated only when the variables are meaningful. For example, you may expect a large positive correlation between
$x=$ number of TV sets per capita and
$y=$ average life expectancy.
However, these variables are not directly related to each other and the correlation coefficient is meaningless.

Read: examples 11.22 (P.215); 11.23 and 11.24 (P.216) and 11.25 (P.217).

Example: Soil temperature $\left(x_{i}\right.$, in $\left.{ }^{0} \mathrm{C}\right)$ and germination interval ( $y_{i}$, in days) were observed for winter wheat at 10 localities:

| $x_{i}$ | 12.5 | 5.0 | 3.0 | 5.0 | 6.5 | 6.0 | 4.0 | 7.0 | 5.5 | 4.0 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $y_{i}$ | 10 | 26 | 41 | 29 | 27 | 19 | 18 | 20 | 28 | 33 |

Find the correlation coefficient between $x$ and $y$.
Solution: We have $n=10$. It is easy to obtain:

$$
\sum_{i=1}^{n} x_{i}=58.5, \sum_{i=1}^{n} y_{i}=251, \sum_{i=1}^{n} x_{i}^{2}=404.75
$$

$$
\begin{gathered}
\sum_{i=1}^{n} y_{i}^{2}=6985, \sum_{i=1}^{n} x_{i} y_{i}=1310.5 \\
L_{x x}=\sum_{i=1}^{n} x_{i}^{2}-\frac{\left(\sum_{i=1}^{n} x_{i}\right)^{2}}{n}=404.75-\frac{58.5^{2}}{10}=62.525 \\
L_{y y}=\sum_{i=1}^{n} y_{i}^{2}-\frac{\left(\sum_{i=1}^{n} y_{i}\right)^{2}}{n}=6985-\frac{251^{2}}{10}=684.9 \\
L_{x y}=\sum_{i=1}^{n} x_{i} y_{i}-\frac{\left(\sum_{i=1}^{n} x_{i}\right)\left(\sum_{i=1}^{n} y_{i}\right)}{n}=1310.5-\frac{(58.5)(251)}{10}=-157.85
\end{gathered}
$$

Now the sample correlation coefficient is

$$
r=\frac{L_{x y}}{\sqrt{L_{x x} L_{y y}}}=
$$

The scatter plot is:
Germination interval on soil temperature


## Notes:

1. The two variables are negatively correlated and $r$ is fairly close to -1 .
2. $r^{2}=$ $\qquad$ .

Therefore, a linear relationship between $x$ (independent variable) and $y$ (dependent variable) will explain about $58 \%$ of the variability in $y$.

Now we consider the fitting of a straight line for a given set of such data with large $|r|$ values. This problem is known as the Simple Linear Regression.

### 12.4 Simple Linear Regression (P.191-198)

Suppose that we have $n$ pairs of observations,

$$
\left(x_{1}, y_{1}\right),\left(x_{2}, y_{2}\right), \cdots,\left(x_{n}, y_{n}\right)
$$

and the scatter plot indicates that there is a possible straight line fit for these data. Now we wish to find (or fit) a straight line to these points.

To find a straight line we need the slope (or the gradient) and the $y$-intercept. Recall the following from your year 8-12 work:

- Equation of a straight line is $y=\alpha+\beta x$.
- $\beta$ is the slope and $\alpha$ is the $y$ intercept.


## Diagram



The regression line

### 12.4.1 Estimates of $\alpha$ and $\beta$

Let $a$ and $b$ be the estimates of the true parameters $\alpha$ ( $y$-intercept) and $\beta$ (slope). It is known that

$$
b=\frac{L_{x y}}{L_{x x}}
$$

is an estimator of $\beta$ and

$$
a=\bar{y}-b \bar{x}
$$

is an estimator of $\alpha$.
Regression line: The estimated regression (or fitted) line is given by

$$
\hat{y}=a+b x
$$

## Interpretation of the regression slope:

For every additional unit increase in $x$, the average change in $y$ is $b$ units.

Note:

1. $a$ and $b$ are called the least squares estimators, because they minimise the sum of squared distances between observed $y_{i}$ 's and estimated $\hat{y}_{i}$ 's.
2. These formulae are also available on the formulae sheet.

Read: Examples 11.8 to 11.11 on P.194-195.

Example: The dose $x$ (in gms) and concentration in urine, $y$ (in $\mathrm{mg} / \mathrm{gm}$ ) of a fluid after intravenous administration are measured for 12 people:

| $x:$ | 46 | 53 | 37 | 42 | 34 | 29 | 60 | 44 | 41 | 48 | 33 | 40 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $y:$ | 12 | 14 | 11 | 13 | 10 | 8 | 17 | 12 | 10 | 15 | 9 | 13 |

(i) Prepare a scatter plot for these data. What do you notice?
(ii) Find the correlation coefficient between $x$ and $y$. What proportion of variability is explained by a simple linear regression model?
(iii) Fit the least squares regression for these data.
(iv) Estimate the urine concentration when the dose is 50 g .

## Solution:

(i)


The scatter plot shows a positive relationship between $x$ and $y$.
(ii) We have $n=12$ and

$$
\begin{aligned}
\sum_{i} x_{i} & =507 \Rightarrow \bar{x}=\frac{507}{12}=42.25 \\
\sum_{i} y_{i} & =144 \Rightarrow \bar{y}=\frac{144}{12}=12 \\
\sum_{i} x_{i}^{2} & =22265, \sum_{i} y_{i}^{2}=1802, \sum_{i} x_{i} y_{i}=6314 \\
L_{x x} & =\sum_{i} x_{i}^{2}-\frac{\left(\sum_{i} x_{i}\right)^{2}}{n}=\underline{22265-\frac{507^{2}}{12}=844.25} \\
L_{y y} & =\sum_{i} y_{i}^{2}-\frac{\left(\sum_{i} y_{i}\right)^{2}}{n}=\underline{1802-\frac{144^{2}}{12}=74} \\
L_{x y} & =\sum_{i} x_{i} y_{i}-\frac{\left(\sum_{i} x_{i}\right)\left(\sum_{i} y_{i}\right)}{n}=\underline{6314-\frac{507(144)}{12}}=230
\end{aligned}
$$

Correlation coefficient:

$$
\begin{gathered}
r=\frac{L_{x y}}{\sqrt{L_{x x} L_{y y}}}= \\
r^{2}= \\
\end{gathered}
$$

and therefore a simple linear regression model explains about $85 \%$ of the variability in $y$.
(iii) Regression line:

$$
b=\frac{L_{x y}}{L_{x x}}=
$$

$$
a=\bar{y}-b \bar{x}=
$$

$\qquad$ .

Fitted line: $\qquad$
or $\qquad$ .

Hence for each additionally administered gram of fluid the concentration in urine increases by $0.272 \mathrm{mg} / \mathrm{gm}$.
(iv) When the dose is 50 g , put $x_{0}=50$. We have

$$
\hat{y}_{0}=a+b x_{0}=
$$

$\qquad$ .

The following graph gives the fitted line plot with fitted value $\hat{y}$ for this example.

Fitted line plot


Note that $\hat{y}_{0}$ provides a good prediction only when $x_{0}$ lies within the range of $x$, i.e. 33 to 60 in this example, or at least close to the range of $x$.

### 12.5 Assessing the Goodness of Fit of a Regression Line

The value of $r^{2}$ is often reported as a measure of the overall goodness of fit of the regression model. In other words, the closer the value of $r^{2}$ is to 1 , the better is the model fit. Notice that perfect straight lines with $r=-1$ or 1 , result in perfect $r^{2}=1$.

### 12.5.1 Residuals

The regression line is a mathematical model for the overall pattern of a linear relationship between $x$ and $y$. The deviations from this overall pattern are called residuals.

Definition: A residual is the difference between an observed value and its corresponding predicted value:

$$
\text { Residual }=y-\hat{y}
$$

Example: Compute the residuals from the dose and concentration data discussed before.

Solution: The fitted line for this data is $\hat{y}=0.490+0.272 x$. We first find $\hat{y}$ for each value of $x$. For example, at $x_{1}=46$,

$$
\hat{y}_{1}=a+b x_{1}=
$$

$\qquad$

Therefore the residual (or error) at $x_{1}=46$ is
$r_{1}=y_{1}-\hat{y}_{1}=$ $\qquad$ .

Similarly, at $x_{2}=53$, the residual is
$r_{2}=y_{2}-\hat{y}_{2}=$ $\qquad$

Fitted line plot


Exercise: Find the remaining 10 residuals for the above regression.

Ans: $\quad 0.446,1.086,0.269,-0.378,0.190,-0.458,-1.642,1.454$, -0.466, 1.630

### 12.5.2 Residual plots

Scatter plots of the residuals $e$ vs. predicted $\hat{y}$ or vs. the $x$ variable can be used to assess the goodness of fit of the regression line.

Example 1: Dose vs. concentration


Example 2: Temperature vs. germination time


Note: The fit is good if the dots are a random scatter around zero.

